Multiplicative Surrogate Standard Deviation – A Group Metric for the Glycemic Variability of Individual Hospitalized Patients

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ABBREVIATIONS USED:

BG = BLOOD GLUCOSE
SD = STANDARD DEVIATION
MSD = MULTIPLICATIVE STANDARD DEVIATION
GGM = GROUP GEOMETRIC MEAN
MSSD = MULTIPLICATIVE SURROGATE STANDARD DEVIATION

Key Words:

Glycemic variability .
Hyperglycemia .
Hypoglycemia .
Critical illness .

Figure count: one figure

Table count: one table, and one text box
ABSTRACT

Objective: Group metrics are described to quantify blood glucose (BG) variability of hospitalized patients.

Methods: The “multiplicative surrogate standard deviation” (MSSD) is the reverse-transformed group mean of the standard deviations (SD) of the logarithmically-transformed blood glucose (BG) data set of each patient. The “geometric group mean” (GGM) is the reverse-transformed group mean of the means of the logarithmically-transformed BG data set of each patient. Before reverse-transformation is performed, the mean of means and mean of SD’s each has its own SD, which becomes a multiplicative standard deviation (MSD) after reverse-transformation. Statistical predictions and comparisons of parametric or nonparametric tests remain valid after reverse-transformation. A subset of a previously-published BG data set of 20 critically ill patients from the first 72 hr of treatment under the SPRINT protocol was transformed logarithmically. After rank-ordering according to the mean of the SD of the logarithmically-transformed BG data of each patient, the cohort was divided into 2 equal groups, those having lower or higher variability.

Results: For the entire cohort, the GGM was 106 mg/dL (÷/× 1.07), and MSSD was 1.24 (÷/× 1.07). For the subgroups having lower and higher variability respectively, the GGM in mg/dL did not differ, 104 (÷/× 1.07) vs. 109 (÷/× 1.07), but the MSSD differed, 1.17 (÷/× 1.03) vs. 1.31 (÷/× 1.05), p = 0.00004.

Conclusion: By using the MSSD with its MSD, groups can be characterized and compared according to glycemic variability of individual patient members.
INTRODUCTION

Observational data suggest that across a wide variety of settings and medical conditions, and perhaps independently of overall glycemia or hypoglycemia, the outcomes of hospitalized patients may be associated with glycemic variability. Our understanding of the impact of glycemic variability has been hampered in some studies by failure to apply variability metrics separately to the blood glucose (BG) distribution of each patient prior to analysis of group characteristics. Observational studies examining central tendency and dispersion of BG of each patient have been hampered by timing of blood glucose tests at sporadic intervals. Additionally, we lack methodologies for controlling variability, by which a randomized controlled trial might be attempted. In this paper we address an additional barrier, the lack of consensus on an appropriate metric for glycemic variability for hospitalized patients, by proposing use of a group metric that permit quantitative description and group comparisons of glycemic variability experienced by individual group members.

To study the impact of variability upon patient outcomes, it is important to recognize and quantitate patient level glycemic variability of individuals and subgroups. In a number of studies, the standard deviation (SD) of the BG of the individual patient has been shown to correlate with hospital outcomes [1-6]. A major time-dependent change of overall glycemia could alter the dispersion of BG in relation to the overall mean BG on the time interval during which the sampling occurred, increasing the SD without necessarily signifying a pattern of recurring large oscillations [7]. It is acknowledged that infrequent sampling results in missing of peaks and nadirs of BG. However, if the moving average of BG is relatively stable, and if timing of sampling is consistent, then SD may mirror the relative amplitude of typical glycemic excursions among patients within a group. Therefore, many authors consider the SD to be a variability metric. Despite its record of performance as a predictor of outcomes, nevertheless the SD is misapplied when used for data sets that are not normally distributed. Population and individual patient BG distribution data typically are positively skewed [8-10], such that use of SD for untransformed data is not appropriate for description of dispersion or for utilization of parametric statistical tests that assume a normal distribution of data. Studies reporting SD can yield results that could predict among patients in the lowest BG range that some BG values would be less than zero [2].

BG data often are capable of being normalized by logarithmic transformation [9-11]. If original data generally has a distribution that is close to log-normal, then the purpose of the logarithmic transform is to gain the advantages of representing the same data points as members of a normal distribution. The BG distribution characteristically seen has a positive skew (a long tail to the right), and with mean greater than median. One advantage of performing logarithmic transformation is to put the data into a symmetrical form in which the mean approximates the median, and the calculated SD creates specific expectations under an empirical rule, predicting the percentages of measurements falling symmetrically within 1 or 2 SD of the mean. Group metrics can be performed on the logarithmically transformed data. If the transformation creates a normally distributed data set, then, assuming other conditions are met (such as independence of observations), between-group analyses using parametric tests are potentially valid. Reverse transformation serves the purpose of returning values that are in units of measure familiar to the reader, on the same scale as the original data. A standard deviation which is added or subtracted in “log space” to give interval bounds becomes a multiplicative SD after reverse
transformation. Interval bounds are given as a mean (± SD) in “log space” or, after reverse transformation, in more familiar units the interval bounds are given as mean (÷ / x multiplicative SD).

The purpose of this report is to describe a process for computing a specific descriptive group metric for glycemic variability experienced by individual patients, which we will call a “multiplicative surrogate SD” of the blood glucose or MSSD. Use of such a metric has been suggested previously. Here we wish to describe its multiplicative characteristics and the details of its application when used together with an artificial geometric mean, which would represent values characteristic of a group member of a cohort (an artificial patient). Since we believe the need for descriptive nomenclature has been a barrier to development of appropriate metrics, we suggest the names “geometrical group mean” (GGM) and “multiplicative surrogate standard deviation” (MSSD) to succinctly denote the metrics described below (Figure 1).
METHODS

Description of method

The purpose is to describe a variability metric representing dispersion of BG values of a typical single artificial patient, as a characteristic of the group to which the patient belongs. In brief, the blood glucose data of each group member is transformed logarithmically. The mean and the SD of the logarithmically-transformed blood glucose data are determined for each patient in the group. For the collection of the glucose means of the logarithmically-transformed data of each patient, a group mean is computed, with its own SD, “in log space.” Using the logarithmically-transformed glucose values, for the collection of SDs for each patient a group mean is computed, with its own SD, again “in log space.” If groups of patients are to be compared using parametric or nonparametric tests, the comparisons should be performed “in log space” (Figure 1).

Reverse-transformation is accomplished using the same base of the logarithm that was used during the initial logarithmic transformation of the BG values. The group mean of the patient means of logarithmically-transformed BG is reverse-transformed to give the GGM. After reverse-transformation, the interval bounds of the 1st and 2nd SD will be asymmetric about the GGM. Reverse-transformation “from log space” of (mean − SD, mean + SD) and (mean − 2 SD, mean +2 SD) gives interval bounds for GGM having the same predictive value as the 1st and 2nd SD determined “in log space” prior to reverse-transformation, i.e. that 67% of values for the population would fall within the 1st SD and 95% within the 2nd SD. Alternatively, if the mean and its SD are reverse-transformed “from log space” separately, the reverse-transformed SD becomes a multiplicative standard deviation (MSD) for the GGM. The same values for interval bounds associated with the 1st and 2nd SD then are given after reverse-transformation by (GGM ÷ MSD, GGM × MSD) and (GGM ÷ MSD², GGM × MSD²). The GGM is a group metric for central tendency of the patient BG.

The group mean of the patient SDs of logarithmically-transformed BG is reverse-transformed to give the MSSD. The interval bounds associated with the 1st and 2nd SD prior to reverse-transformation “from log space” then are given by reverse-transformation of each of the interval bounds (mean − SD), (mean + SD) and (mean − 2 SD), (mean +2 SD). Alternatively, if the mean and its SD are reverse-transformed “from log space” separately, the same values for interval bounds associated with the 1st and 2nd SD then are given after reverse-transformation by (MSSD ÷ MSD, MSSD × MSD) and (MSSD ÷ MSD², MSSD × MSD²). The MSSD is a group metric for variability of the patient BG.

A true SD is defined only in reference to a single mean. The metric MSSD, derived after averaging the SD’s of transformed BG’s from multiple patients “in log space,” is not paired with any specific single mean, and therefore is not properly a SD. Therefore, we suggest rather that the name should imply that that metric is a “surrogate” for a SD. This “SD surrogate” resembles a multiplicative SD. The surrogate SD is unit-less. This “SD surrogate” has magnitude characteristic of a multiplicative SD. When applied to a geometric mean BG that would be found in a BG distribution typical for a member of the sampled population, the MSSD yields values for interval ranges comparable in magnitude to an actual SD. If the MSSD is used in reference to GGM, it is used as follows:
Method applied to demonstration data set

Previously published data will be used to demonstrate the method of analysis. The data set is found at http://www.journalofdst.org/Journal/pdf/July2008/VOL-2-4-ORG4-CHASE-DATA-SUPPLEMENT-DS1.XLS

For the present, a subset of BG data from time zero to 72 hr inclusive was examined for each patient in the cohort reported by Chase and colleagues, a time interval chosen because each patient continued to have data beyond 72 hr but also judged brief enough to capture differences in initial glycemic variability between patients. Some patients experienced brief gaps in data or compliance noted, but none were removed from treatment under the SPRINT algorithm for longer than 2 hr during the interval of data collection.

The BG data set of each patient was logarithmically-transformed. The SDs of the logarithmically-transformed BG data sets were rank-ordered, and the 20 patient members of the cohort were divided into two groups, having SD’s of the transformed BG data set of each patient that were either below or above a value between the two median SD values, i.e. into “lower variability” and “higher variability” groups, each having 10 members.

The term “overall” when applied to mean or SD of a BG distribution will refer to the application of the metric to the set of all eligible BG data of the cohort (or a subgroup), using the BG as the unit of observation. For the entire cohort, and for the lower and higher variability patient subgroups separately, metrics were described in each of 4 ways (Table 1): (1) metrics from untransformed overall BG data; (2) metrics from untransformed BGs of each patient; (3) reverse-transformed metrics from logarithmically-transformed overall BG data; and (4) reverse-transformed metrics from logarithmically-transformed BG data of each patient. As final steps, reverse-transformed measures of central tendency, arithmetic SD’s, and interval bounds were converted to SI units, and results were rounded so as to have no decimal places. After use for other calculations, MSD’s and MSSD in a final step were rounded to 2 decimal places for presentation as results.

The numbers of untransformed BG values were counted and proportions were determined that lay within ±1 SD or ±2 SD or outside of 2 SD of the mean of the untransformed and logarithmically transformed overall BG data, and that lay within 1 MSSD or 2 MSSD of GGM or outside of 2 MSSD of GGM, for the entire cohort and separately for the lower and higher variability subgroups.

Using the unpaired two-tailed T-test with samples having unequal variance, first the overall means of the logarithmically-transformed overall BG data between the lower-variability and the higher-variability patient subgroups were compared, and then the patient means and SD’s of the two groups were compared.
RESULTS

The frequency distribution of BG values is shown for the overall BG values of each of the two groups of 10 patient in Figure 2. The results of 4 methods of analysis of each group are shown in Table 1 for the entire cohort and for the lower and higher variability patient subgroups. In methods (3) and (4) the initial logarithmic transformation converts the data into the form used for development of means and SD’s and for statistical testing, after which, in each method, the means and interval bounds are reverse-transformed monotonically to yield the results shown.

When the patient subgroups with lower and higher variability were compared with respect to their logarithmically-transformed overall BG datasets, consisting of 527 measurements from the lower variability subgroup and 505 measurements from the higher variability subgroup of patients, the mean overall value of the logarithmically transformed BG of the two subgroups differed (p = 0.0027). In method (3), the overall means ± SD of the logarithmically transformed BGs for the lower and higher variability subgroups were 0.760562663 ± 0.072367226 vs. 0.779476971 ± 0.122302184 respectively, which are reverse-transformed to means and MSD’s of 104 (± 1.18) vs. 108 (± 1.33) mg/dL. When the 10 means and 10 SD’s from the logarithmically-transformed BG data sets of all patients within each of the two subgroups were compared, the means of the means did not differ (p = 0.16), but the means of the SD’s differed, p < 0.00004. After reverse-transformation, the GGM’s with MSD were 104 (± 1.07) vs. 109 (± 1.07) mg/dL for the patient groups having the lower and higher variability respectively. Expressed in unit less numbers, the corresponding MSSD’s with MSD’s were 1.17 (± 1.03) vs. 1.31 (± 1.05). Under methods (1), (3) and (4), the actual counts and percentage of BG results from the sampled groups falling within the range of interval bounds are stated for comparison with the statistical prediction for the population. For each of the two groups of 10 patients, the percent of overall BG values for the entire group falling within interval bounds is shown graphically for methods (1), (2), (3), and (4) in Figure 3.

DISCUSSION

The performance of several metrics for evaluation of glycemic variability in the hospital has recently been reviewed. Here we are focusing on proposed improvements to the use of SD as a variability metric. First, it is important to ensure stability of a measure of central tendency during the time of observation, if SD is intended to reflect variability. Second, in this discussion, we address a solution to the problem of the distribution of BG, which characteristically requires transformation for appropriate use of SD for descriptive purposes or performance of parametric statistical comparisons. Third, we focus on the importance of using the patient as the unit of observation for BG metrics, rather than the BG.

The authors do not advocate either of the approaches (1) or (2) shown in Table 1, but display the results for comparison, recognizing that historically important publications on the subject of glycemic variability have used untransformed raw BG data. Although glycemic variability was not the principal focus of the article, the first method, using the BG as the unit of observation, was employed in the Leuven, Belgium 2001 trial of intensive glycemic control in the surgical ICU. The second method of analysis was used by two pivotal observational reports about the interaction between glycemic
variability and outcomes. By using the raw data of the patient as the unit of observation, it was possible to correlate SD with individual outcomes.

Reverse-transformation will monotonically preserve the order of BG values and the values of the interval bounds. As is true in general for multiplicative SD for positively skewed data, after logarithmic transformation the interval bounds are asymmetric about the measure of central tendency, with a wider range above than below. For an excellent visual depiction, the reader is directed to the two-panel Figure 3 in the paper by Limpert et al., which shows the distribution of idealized hypothetical data, drawn as a continuous curve. Using discrete patient data, a graphical comparative display of the untransformed and transformed BG distribution of sample cases has been published by Palerm and colleagues. The effect of transformation is to compress the long right-hand end of the curve, so that the peak of the redrawn curve after logarithmic transformation is centered on the mean of the transformed data. In the present study, as is true in general for geometric means of positively skewed data, in comparison to arithmetic means, the reverse-transformed metrics for central tendency shown in methods (3) and (4) are lower than the arithmetic means shown in methods (1) and (2). The positive skew of untransformed BG data is more apparent in the group having higher variability (Figure 2).

Reverse-transformed metrics may improve the predictive credibility of the interval bounds for overall group BG data. In the present study, under methods (3) and (4), the mean and SD of logarithmically-transformed BG data were used to establish interval bounds. Compared to use of metrics based on untransformed data under methods (1) and (2), the improvement of symmetry of BG measurements about the mean, and the actual percentage distribution of BG measurements between the values demarcated by the mean, 1st and 2nd interval bounds, suggest that the distribution of the logarithmically-transformed BG's has approached a normal distribution (Figure 3).

A limitation of the method of use of logarithmic transformation, reverse transformation, and multiplicative SD is the effect of rounding. The final results here are expressed to 2 decimal places for the multiplicative SD, and none for BG in mg/dL. A limitation of the present study is that it has not been convincingly demonstrated from a large sample of patients whether or not the collection of means or SD's of transformed BG's of each patient would yield a normal distribution of these means or SD's. In case future evaluation shows deviation from a normal distribution, non-parametric testing could be used to compare distributions of these means or SD's.

From the complete data set of the cohort of ICU patients published by Chase and colleagues, the reported overall glycemic average of reverse-transformed BG data was 105 mg/dL, with a first multiplicative SD of 1.2x, yielding a predicted 66% one-SD range of 86-126 mg/dL and a 95% two-SD range of 72-151 mg/dL. Using method (3) in Table 1, we performed a similar analysis of overall results for the same cohort, using logarithmically-transformed BGs, but restricting the data to the first 72 hr of treatment. The overall mean BG after reverse-transformation was 106.1 mg/dL with a one-SD range of 84.2-133.6 mg/dL and a 2-SD range of 66.8-168.4 mg/dL.

Taken together, in method (4) the GGM and MSSD describe an artificial patient, who may be seen as a typical member of a group. The choice of the term “artificial” results from the use of a...
surrogate value for the MSSD and the creation of a GGM, which is actually a reverse-transformed mean of means. The principal advantage of method (4) is that the status of individual patients according to variability becomes evaluable. Going forward, in the field of glycemic control, we need to become sensitized to recognizing ranges of multiplicative SD values that are low or high, without mentally applying them to a mean BG. However, the reverse-transformation permits expression of the GGM and interval bounds in familiar units of measure. The MSSD and GGM are not mathematically linked in the same obligatory manner as a true SD is linked to the distribution of BGs associated with its mean. It cannot be predicted that 67% of measurements from the randomly sampled population will fall within 1 MSSD of the GGM or 95% within 2 MSSD or the GGM. If all patients in the population had the statistics of the BG distribution described for the artificial patient (or if there was only one patient), then 67% of BG values overall for the population would fall within the first interval bounds, and 95% within the second interval bounds.

Use of parametric testing dependent upon mean and SD for group comparisons requires that each group have an approximately normal distribution. Differences between groups in the protocol-controlled population reported here were not demonstrable prior to logarithmic transformation of data. After logarithmic transformation of BGs, using method (3) and method (4), it was possible to identify subgroups with different variability, shown in Table 1 as the middle and right-hand columns. A difference of the overall mean BG was demonstrable statistically only when using method (3) to compare the 527 vs. 505 overall BG values of the two groups of 10 patients each.

The demonstration data patient group was tightly controlled under the SPRINT protocol, having a single target range. One might envision a different situation in which algorithm designs or institutional protocols were capable of aiming at more than one target range, having for example one default target range for general critical care and a second target range for diabetic ketoacidosis. The overall BG distribution after reaching target range control then might exhibit a bimodal pattern, with overall mean intermediate between the two targets. Ideally the mean BG values for each patient would cluster into the differing target ranges, appropriate to the conditions of the patients. The overall SD of BG’s considered collectively for the group might overestimate the variability experienced by individual patients. The SD for the BG of each patient then need not be a high value, but might be proportionate to the mean BG achieved, probably with similar coefficient of variability (CV) between patients. These two theoretical examples -- (a) tight control of central tendency despite differences of individual variability, and (b) differences of central tendency despite tight control of individual variability -- are used to emphasize the potential importance of patient-level determination of group metrics for central tendency and variability.

It is proposed that the GGM may be used to study the relationship of overall glycemia to outcomes and also the effectiveness of algorithms in achieving desired targets for groups or subgroups of patients. The GGM describes the central tendency characteristic of patient members of a group, and, although the GGM entails logarithmic transformation, by reverse-transformation GGM expresses the central tendency metric in units familiar to the reader.
There is a need to examine the impact of variability upon the outcomes of individual patients. In order to do so, a metric is required by which the variability of a patient can be compared to the typical variability that is characteristic of a patient member of his or her group. In particular, as variability is associated with individual patient outcomes, then there is a need to present patient-specific variability metrics, but also, as presented here, to group those metrics in a concise and useful way in presenting and analyzing larger studies. We propose that the MSSD is a candidate metric expressed in familiar units of measure that may be used to described the typical variability of an individual group or subgroup member, for study of relationship of patient variability to outcomes.

The proposed metrics GGM and MSSD have been developed and validated in a very small set of patients. We should seek to evaluate variability as a predictor of outcomes, independent of hypoglycemia or severe hyperglycemia. Assuming prevention of hypoglycemia can be achieved, it may be argued that we do not know the relative burden of medical strategies that might minimize variability, as compared to any burden resulting from glycemic variability itself. Improvements in insulin algorithms and the development of non-insulin-based strategies may permit future studies to be conducted that may randomize patients to greater or lesser glycemic variability without significant differences in hypoglycemia. Evaluation eventually will be required to examine the ability of GGM and MSSD, compared to other metrics for variability and central tendency, to predict nonglycemic outcomes.

CONCLUSIONS

The GGM and MSSD are presented as group metrics, requiring logarithmic transformation of the BG data set of each patient. Development of statistics before a final reverse-transformation permits identification of predictive interval bounds and application of statistical testing for group comparisons.
FUNDING SOURCES: none

ACKNOWLEDGEMENTS: none

DISCLOSURES: no conflict of interest related to this work
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<tr>
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<th>Entire cohort</th>
<th>Lower variability subgroup a</th>
<th>Higher variability subgroup a</th>
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</thead>
<tbody>
<tr>
<td>Number of patients</td>
<td>20</td>
<td>10</td>
<td>10</td>
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<tr>
<td>Number of BG results</td>
<td>1032</td>
<td>527</td>
<td>505</td>
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(1a) Metrics from untransformed overall BG data  
(n = number of BG results)

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<th>Entire cohort</th>
<th>Lower variability subgroup a</th>
<th>Higher variability subgroup a</th>
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</thead>
<tbody>
<tr>
<td>Overall mean in mg/dL</td>
<td>109 (± 28)</td>
<td>105 (± 17)</td>
<td>113 (± 35)</td>
</tr>
<tr>
<td>Interval bounds in mg/dL, ± 1 SD</td>
<td>(82 – 137)</td>
<td>(88 – 123)</td>
<td>(78 – 148)</td>
</tr>
<tr>
<td>Interval bounds in mg/dL, ± 2 SD</td>
<td>(54 – 164)</td>
<td>(71 – 140)</td>
<td>(43 – 183)</td>
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(1b) Distribution of BG’s

<table>
<thead>
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<th>Entire cohort</th>
<th>Lower variability subgroup a</th>
<th>Higher variability subgroup a</th>
</tr>
</thead>
<tbody>
<tr>
<td>Number (%) of BG within mean ± 1 SD’s b</td>
<td>814 (78.9%)</td>
<td>384 (72.9%)</td>
<td>394 (78.0%)</td>
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<tr>
<td>Number (%) of BG within mean ± 2 SD’s b</td>
<td>986 (95.5%)</td>
<td>499 (94.7%)</td>
<td>486 (96.2%)</td>
</tr>
</tbody>
</table>

(2) Metrics from untransformed BGs of each patient in mg/dL  
(n = number of patients)

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<tr>
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<th>Entire cohort</th>
<th>Lower variability subgroup a</th>
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</thead>
<tbody>
<tr>
<td>Mean of BG means of each patient</td>
<td>109</td>
<td>106</td>
<td>113</td>
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<tr>
<td>Mean of SD’s of each patient</td>
<td>24</td>
<td>16</td>
<td>32</td>
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(3a) Reverse-transformed metrics from logarithmically-transformed overall BG data  
(n = number of BG results)

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<th>Lower variability subgroup a</th>
<th>Higher variability subgroup a</th>
</tr>
</thead>
<tbody>
<tr>
<td>Overall mean BG in mg/dL (÷ / × overall MSD)</td>
<td>106 (÷ / × 1.26)</td>
<td>104 c (÷ / × 1.18)</td>
<td>108 c (÷ / × 1.33)</td>
</tr>
<tr>
<td>Interval bounds in mg/dL within 1st MSD</td>
<td>(84 – 134)</td>
<td>(88 – 123)</td>
<td>(82 – 144)</td>
</tr>
<tr>
<td>Interval bounds in mg/dL within 2nd MSD</td>
<td>(67 – 168)</td>
<td>(74 – 145)</td>
<td>(62 – 191)</td>
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(3b) Distribution of BG’s

<table>
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<th>Entire cohort</th>
<th>Lower variability subgroup a</th>
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<tbody>
<tr>
<td>Number (%) of BG within 1st MSD</td>
<td>788 (76.4%)</td>
<td>384 (72.9%)</td>
<td>370 (73.3%)</td>
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<tr>
<td>Number (%) of BG within 2nd MSD</td>
<td>975 (94.5%)</td>
<td>499 (94.7%)</td>
<td>480 (95.0%)</td>
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</table>

(4a) Reverse-transformed metrics from logarithmically-transformed BG data of each patient  
(n = number of patients)

<table>
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<th>Entire cohort</th>
<th>Lower variability subgroup a</th>
<th>Higher variability subgroup a</th>
</tr>
</thead>
<tbody>
<tr>
<td>GGM in mg/dL (÷ / * 1 MSD of GGM)</td>
<td>106 (÷ / × 1.07)</td>
<td>104 (÷ / × 1.07)</td>
<td>109 (÷ / × 1.07)</td>
</tr>
<tr>
<td>MSSD (÷ / * 1 MSD of MSSD)</td>
<td>1.24 (÷ / × 1.07)</td>
<td>1.17 d (÷ / × 1.03)</td>
<td>1.31 d (÷ / × 1.05)</td>
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<tr>
<td>GGM in mg/dL, ÷ / * MSSD once</td>
<td>(86 – 132)</td>
<td>(89 – 122)</td>
<td>(83 – 142)</td>
</tr>
<tr>
<td>GGM in mg/dL, ÷ / * MSSD twice</td>
<td>(70 – 163)</td>
<td>(76 – 142)</td>
<td>(64 – 186)</td>
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(4b) Distribution of BG’s

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<th>Higher variability subgroup a</th>
</tr>
</thead>
<tbody>
<tr>
<td>Number (%) of BG within 1 MSSD of GGM*</td>
<td>754 (73.1%)</td>
<td>360 (68.3%)</td>
<td>360 (71.3%)</td>
</tr>
<tr>
<td>Number (%) of BG within 2 MSSD of GGM*</td>
<td>959 (92.9%)</td>
<td>491 (93.2%)</td>
<td>478 (94.7%)</td>
</tr>
</tbody>
</table>
Footnotes to Table 1.

BG = blood glucose
SD = standard deviation
MSD = multiplicative standard deviation
GGM = group geometric mean
MSSD = multiplicative surrogate standard deviation

a The SDs of the logarithmically-transformed BG data sets were rank-ordered, and the patient members of the cohort were divided into lower and upper halves (see text)
b The actual counts of BG results from the sampled groups are given
c $p = 0.0027$.
d $p < 0.00004$
Legend for Figure 1.

A process is depicted for calculation of geometric group mean (GGM) and multiplicative surrogate standard deviation (MSSD), intended to permit characterization and comparison of groups of hospitalized patients according to glycemic variability of patient members of each group. In case the collection of means or SD’s of transformed BG’s of each patient does not yield a normal distribution of these means or SD’s, instead of using the standard deviation (SD), it would be appropriate to consider interquartile ranges and to use non-parametric testing for comparisons.
Legend Figure 2. The overall frequency distribution for each of two patient groups (n=10 in each group) described in the text is shown as the number of blood glucose (BG) measurements falling within bins of BG concentration incrementally increasing from left to right by 10 mg/dL between markers.
Figure 3. Interval bounds are established under each of four methods with reference to the following metrics: the mean and SD using untransformed overall BG data (method 1); the mean of BG means and the mean of SD’s of each patient, using untransformed BGs of each patient (method 2); the overall geometric mean and multiplicative SD (MSD) using reverse-transformed metrics from logarithmically-transformed overall BG data (method 3); and the reverse-transformed mean of means (group geometric mean, GGM) and mean of SD’s (multiplicative surrogate standard deviation, MSSD) of logarithmically-transformed BG data of each patient (method 4). Interval bounds (< -2; -2 to -1; -1 to mean; mean to +1; +1 to +2; and > +2) refer to values obtained by use of a mean and an arithmetic SD (Panel A), or a geometric mean and multiplicative SD (Panel B), to define ranges bounded by BG values. The interval bounds in methods 1 and 2 equal mean ± 1 or ± 2 SD’s. The interval bounds in method 3 equal geometric mean (± / x MSD) or geometric mean (± / x MSD^2). The interval bounds in method 4 equal...
GGM ($\div / x \text{MSSD}$) or ($\div / x \text{MSSD}^3$). To determine the percent of BG measurements falling within interval bounds, each of the four methods of metrics for central tendency and variability is applied to the overall BG data of two patient groups described in the text (n = 10 patients in each group, having lower and higher variability), to determine percent of BG’s (n=527 and 505 BG measurements in the lower and higher variability groups respectively) within the interval bounds defined under each method.
REFERENCES (PLEASE SEE COVER LETTER ABOUT THE STATUS OF REFERENCE 10, WHICH IS UNDER REVIEW)